

EE 525
Notes on Random Variables
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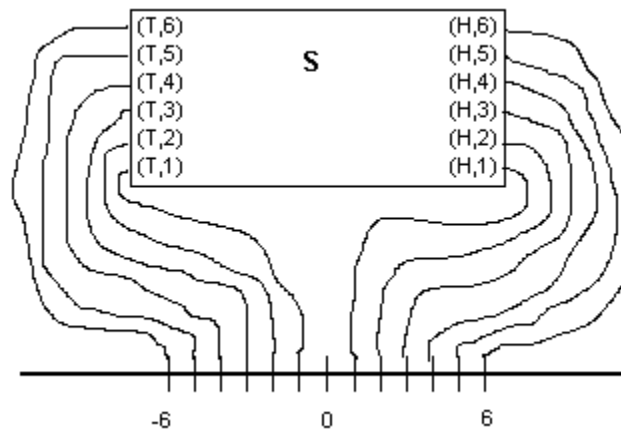
Random Variable

A random variable X can be considered a function that maps all elements of the sample space into points on the real number line or some parts thereof.

Example 1:

An experiment consists of rolling a die and flipping a coin. The random variable is a function X chosen such that:

- (1) a coin head (H) corresponds to positive values of X that are equal to the number shown on the die and
- (2) a coin tail (T) corresponds to negative values of X that are equal to the number shown on the die



Example 2:

A pointer on a wheel is spun. The possible outcomes are the numbers 0 to 12 marked on the wheel. The sample space consists of the numbers in the set $\{0 < s \leq 12\}$. The random variable (r.v.) is defined by: $X = X(s) = s^2$.

An r.v. is a single valued function and the set $\{X \leq x\}$ is an event for any real number.

$$P\{X = -\infty\} = 0, \quad P\{X = \infty\} = 0$$

Example 1 illustrates a discrete r.v. and example 2 illustrates a continuous r.v.

Examples 1 and 2 illustrate the freedom provided in the mapping between the “objects” in a chance experiment (such as coins and dice) and the values of a random variable, on the real number line.

Distribution Function

$P\{X \leq x\}$ is the probability of the event $\{X \leq x\}$. We call this function the “cumulative probability distribution function” of the r.v. X and it is denoted by $F_X(x)$.

$$F_X(x) = P\{X \leq x\}$$

Properties:

1. $F_X(-\infty) = 0$
2. $F_X(\infty) = 1$
3. $0 \leq F_X(x) \leq 1$
4. $F_X(x_1) \leq F_X(x_2)$ if $x_1 < x_2$
5. $P\{x_1 < X \leq x_2\} = F_X(x_2) - F_X(x_1)$
6. $F_X(x^-) = F_X(x)$

If X is a discrete r.v., then

$$F_X(x) = \sum_{i=1}^N P\{X = x_i\} u(x - x_i)$$

where $u(x) = \begin{cases} 1 & x > 0 \\ 0 & x < 0 \end{cases}$ and $P(x_i) = P\{X = x_i\}$

$$\text{i.e., } F_X(x) = \sum_{i=1}^N P(x_i) u(x - x_i)$$

Density Function

$$f_X(x) = \frac{dF_X(x)}{dx}$$

$f_X(x)$ is the “density function” of the r.v. X . For discrete r.v.’s $\delta(x)$ is used to describe the derivative of $F_X(x)$. $\delta(x)$ is described by its integral property:

$$\Phi(x_0) = \int_{-\infty}^{\infty} \Phi(x) \delta(x - x_0)$$

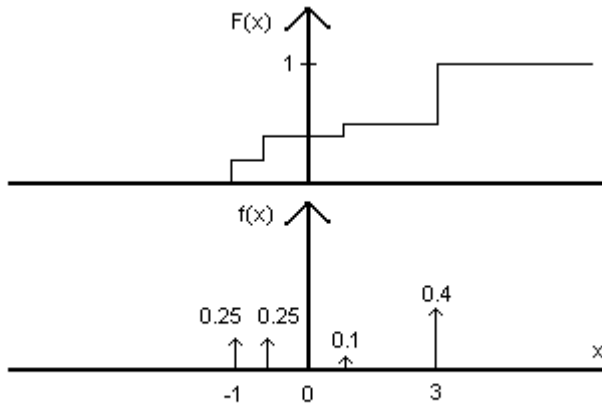
where $\Phi(x)$ is a function continuous at $x = x_0$.

$$\delta(x) = \frac{du(x)}{dx} \text{ or } \int_{-\infty}^x \delta(\zeta) d\zeta = u(x)$$

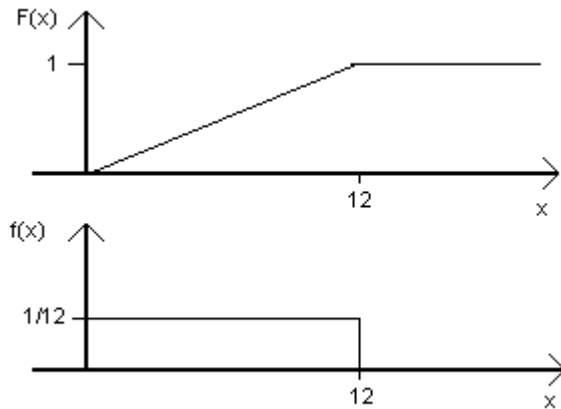
$$f_X(x) = \sum_{i=1}^N P(x_i) \delta(x - x_i)$$

Example:

A discrete r.v.



A continuous r.v.



Properties of $f_X(x)$:

1. $0 \leq f_X(x) \quad \forall x$
2. $\int_{-\infty}^{\infty} f_X(x) dx = 1$
3. $F_X(x) = \int_{-\infty}^x f_X(\zeta) d\zeta$
4. $P\{x_1 < x < x_2\} = \int_{x_1}^{x_2} f_X(x) dx$

Expected Value:

Expected value of an r.v. X is defined by:

$$E[X] = \bar{X} = \int_{-\infty}^{\infty} x f_X(x) dx$$

If X is discrete with N possible values having possible values $P(x_i)$, then

$$f_X(x) = \sum_{i=1}^N P(x_i) \delta(x - x_i) \text{ and } E[X] = \sum_{i=1}^N x_i P(x_i)$$

Expected Value of a Function of a Random Variable

$$E[g(X)] = \int_{-\infty}^{\infty} g(x)f_X(x)dx \text{ or, if } X \text{ is discrete, } E[X] = \sum_{i=1}^N g(x_i)P(x_i)$$

Moments

Moments about the origin have $g(X) = X^n \quad x = 0,1,2,3,\dots$

The n^{th} moment about the origin is given by $m_n = E[X^n] = \int_{-\infty}^{\infty} x^n f_X(x)dx$.

If $n = 1$, $m_1 = E[X] = \bar{X}$

Moments about the mean value of X are called central moments, denoted by μ_n . They are defined as the expected value of the function $g(X) = (X - \bar{X})^n \quad n = 0,1,2,\dots$

$$\mu_n = E[(X - \bar{X})^n] = \int_{-\infty}^{\infty} (x - \bar{X})^n f_X(x)dx$$

$$\mu_0 = 1, \text{ the area under } f_X(x) \quad \mu_1 = E[(X - \bar{X})] = \int_{-\infty}^{\infty} (x - \bar{X})f_X(x)dx = 0$$

Variance

Variance is defined as the second central moment and is denoted by σ_X^2 .

$$\sigma_X^2 = \mu_2 = E[(X - \bar{X})^2] = \int_{-\infty}^{\infty} (x - \bar{X})^2 f_X(x)dx$$

σ_X is called the standard deviation of X .

$$\begin{aligned} \sigma_X^2 &= E[X^2 - 2X\bar{X} + \bar{X}^2] = E[X^2] - 2\bar{X}E[X] + \bar{X}^2 = E[X^2] - 2\bar{X}^2 + \bar{X}^2 \\ &= E[X^2] - \bar{X}^2 = m_2 - m_1^2 \end{aligned}$$

Characteristic Function

The characteristic function of a r.v. X is defined by

$$\Phi_X(\omega) = E[e^{j\omega X}] = \int_{-\infty}^{\infty} f_X(x)e^{j\omega x} dx \quad \text{where } j = \sqrt{-1}$$

$$f_X(x) = \frac{1}{2\pi} \int_{-\infty}^{\infty} \Phi_X(\omega)e^{-j\omega x} d\omega$$

$$m_n = (-j)^n \left. \frac{d^n \Phi_X(\omega)}{d\omega^n} \right|_{\omega=0}$$

Moment Generating Function

$$M_X(\nu) = E[e^{\nu X}] = \int_{-\infty}^{\infty} f_X(x)e^{\nu x} dx \quad \text{where } -\infty < \nu < \infty$$

$$m_n = \left. \frac{d^n M_X(\nu)}{d\nu^n} \right|_{\nu=0}$$

Gaussian or Normal Random Variables

A random variable X is called Gaussian if its probability density function has the form

$$f_X(x) = \frac{1}{\sqrt{2\pi}\sigma_X} e^{-\frac{(x-m_x)^2}{2\sigma_X^2}}$$

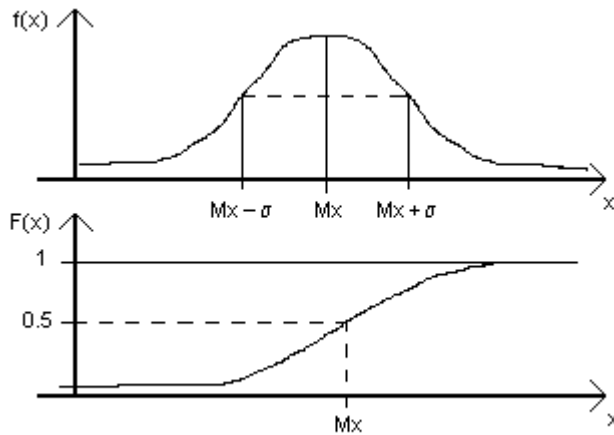
The density function has two parameters m_x and σ_X^2 where

$$m_x = E[X] = \int_{-\infty}^{\infty} x f_X(x) dx \quad \text{and} \quad \sigma_X^2 = \int_{-\infty}^{\infty} (x - m_x)^2 f_X(x) dx$$

For shorthand notation we can write $X \sim N(m_x, \sigma_X^2)$

$$F_X(x) = \frac{1}{\sqrt{2\pi}\sigma_X} \int_{-\infty}^x e^{-\frac{(\zeta-m_x)^2}{2\sigma_X^2}} d\zeta$$

This integral does not have a closed form solution and must be evaluated by numerical methods.



Binomial Distribution

If $0 < p < 1$ and $n = 1, 2, 3, \dots$ then

$$f_X(x) = \sum_{k=0}^n \binom{n}{k} p^k (1-p)^{n-k} \delta(x-k)$$

is called the binomial density function. The binomial distribution function is found:

$$F_X(x) = \sum_{k=0}^n \binom{n}{k} p^k (1-p)^{n-k} u(x-k)$$

Exponential Distribution

$$f_X(x) = \begin{cases} \frac{1}{b} e^{-\frac{x-a}{b}} & x > a \\ 0 & x < a \end{cases}$$

$$F_X(x) = \begin{cases} 1 - e^{-\frac{x-a}{b}} & x > a \\ 0 & x < a \end{cases}$$

Examples

Example 1:

Using the exponential distribution:

$$E[X] = \int_a^{\infty} \frac{x}{b} e^{-\frac{x-a}{b}} dx = \frac{1}{b} e^{\frac{a}{b}} \int_a^{\infty} x e^{-\frac{x}{b}} dx = \frac{1}{b} e^{\frac{a}{b}} e^{\frac{x}{b}} (-bx - b^2) \Big|_a^{\infty} = a + b$$

Example 2:

A r.v. is Rayleigh distributed e.g., $f_X(x) = \begin{cases} \frac{2}{b}(x-a)e^{-\frac{(x-a)^2}{b}} & x > a \\ 0 & x < a \end{cases}$ with $a = 0$, and $b = 5$.

$$Y = g(V) = V^2$$

$$E[g(V)] = E[V^2] = \int_0^{\infty} \frac{2u^3}{5} e^{-\frac{u^2}{5}} du$$

Let $\zeta = \frac{u^2}{5}$, so $d\zeta = \frac{2}{5} u du$ and $E[g(V)] = 5 \int_0^{\infty} \zeta e^{-\zeta} d\zeta = 5$

Example 3:

Using the exponential distribution:

$$\sigma_X^2 = \int_a^{\infty} (x - \bar{X})^2 \frac{1}{b} e^{-\frac{x-a}{b}} dx \quad \text{Now let } u = x - \bar{X}.$$

$$\sigma_X^2 = \frac{1}{b} e^{-\frac{\bar{X}-a}{b}} \int_{a-\bar{X}}^{\infty} u^2 e^{-\frac{u}{b}} du = (a + b - \bar{X})^2 + b^2$$

From Example 1, $\bar{X} = a + b$, so $\sigma_X^2 = b^2$

Example 4:

Using the exponential distribution:

$$\Phi_X(\omega) = \int_a^{\infty} \frac{1}{b} e^{-\frac{x-a}{b}} e^{j\omega x} dx = \frac{e^{\frac{a}{b}}}{b} \int_a^{\infty} e^{-\left(\frac{1}{b} - j\omega\right)x} dx = \frac{1}{b} e^{\frac{a}{b}} \frac{e^{-\left(\frac{1}{b} - j\omega\right)x}}{-\left(\frac{1}{b} - j\omega\right)} \Big|_a^{\infty} = \frac{e^{j\omega a}}{1 - j\omega b}$$

$$\frac{d\Phi_X(x)}{d\omega} = e^{j\omega a} \left[\frac{ja}{1 - j\omega b} + \frac{jb}{(1 - j\omega b)^2} \right]$$

$$m_1 = (-j) \frac{d\Phi_X(\omega)}{d\omega} \Big|_{\omega=0} = a + b$$